A COMPARISON OF THREE CONTROL ALGORITHMS AS
APPLIED TO THE MONETARIST-FISCALIST DEBATE

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1. Introduction

In this paper we shall use an optimal control framework to examine the relative effectiveness of monetary and fiscal policies for the purpose of controlling the major macroeconomic aggregates. In section 2, we shall present a very simple linear macroeconomic model with additive Gaussian disturbances. After briefly describing three control algorithms in section 3, we shall, in section 4, apply these three control algorithms to our simple model. Monetary policy will be represented by the money supply and fiscal policy will be represented by government expenditures. In evaluating the effectiveness of a given instrument, we shall designate that instrument as a discretionary instrument and the other instrument as a passive instrument, and then solve an optimal control problem. The values of the discretionary instrument are determined subject to feedback control, but the values of the passive instrument are constrained to change at a constant rate over the planning horizon. In addition, we solve the control problem with both instruments assumed to be discretionary. Comparison of the expected welfare costs in the three situations serves to evaluate the effectiveness of each discretionary instrument.

Three different algorithms will be used to perform our analysis of the relative effectiveness of monetary and fiscal policies
in order to determine whether our results are sensitive to the choice of control algorithm employed. The three algorithms presented in this paper differ in their treatments of learning. Method I is a certainty equivalence control algorithm formulated by Chow (1972). The assumptions of certainty equivalence preclude the possibility of learning by assuming that the parameters of the linear model are known with certainty. Method II, which is presented by Chow (1973a), recognizes the uncertainties in the parameters of the model but ignores the possibility of learning. Method III, which is a dual adaptive control algorithm presented by Chow (1973c), anticipates that learning will occur through the re-estimation of the unknown parameters of the linear model as additional observations are obtained with the passage of time. Method III is closest to being optimal and contains method I and method II as special cases.

2. A Simple Macroeconomic Model

For the policy analysis of this paper, we shall employ a very simple aggregative model. It is based on real quarterly data covering the period from 1954-I to 1963-IV, which corresponds roughly to the period between the end of the Korean War and the beginning of heavy United States involvement in Vietnam. It consists of only two endogenous target variables, consumption \( (C_t) \) and investment \( (I_t) \), and two instruments, government expenditures \( (E_t) \) and the money supply \( (M_t) \). We assume that in the short-run, government authorities can control \( E_t \) and \( M_t \) in real terms since prices do not change.
rapidly enough to seriously offset their actions. Over the time period covered by our data, the rate of inflation was low enough to make this assumption plausible.

Our model is based on a closed economy. Desired consumption is a linear function of GNP, and the realized period-to-period adjustment in consumption is subject to a partial adjustment factor:

\[ C_t = aC_{t-1} + bI_t + bE_t + d. \]

The structural equation for investment is based upon a modification of Samuelson's private consumption accelerator. We posit that the desired level of the capital stock is a linear function of consumption and that the realized adjustment of the capital stock is subject to a partial adjustment factor. Since gross investment, \( I_t \), is defined as \( K_t - (1-D)K_{t-1} \), where \( D \) is the depreciation rate of the capital stock, we have

\[ I_t = eC_t - (1-D)eC_{t-1} + fI_{t-1} + g. \]

In addition, we assume that the level of gross investment is linearly related to the money supply in order to capture some of the effects of interest rates upon investment:

\[ I_t = e'C_t - (1-D)e'C_{t-1} + f'I_t + hM_t + g'. \]

The estimated reduced form equations corresponding to the structural equations are
(2.4) \[ C_t = 0.9266C_{t-1} - 0.0203I_{t-1} + 0.3190E_t + 0.4206M_t - 63.2386; \]
\[ (0.0534)_{t-1} (0.0916)_{t-1} (0.1389) \]
\[ R^2 = 0.9958 \]
\[ D-W = 1.7084 \]

(2.5) \[ I_t = 0.1527C_{t-1} + 0.3806I_{t-1} - 0.0735E_t + 1.5389M_t - 210.8994; \]
\[ (0.0781)_{t-1} (0.1339)_{t-1} (0.2031) \]
\[ R^2 = 0.8749 \]
\[ D-W = 1.7582 \]

Note that each of these estimated equations has a high value of \( R^2 \). In addition, the Durbin-Watson statistic, although biased toward 2.0 because of the lagged endogenous variable, does not suggest significant serial correlation in either equation.

A criticism that may be raised against the above model is that it includes only the current values of \( M_t \) and \( E_t \) among the explanatory variables. However, concerning the lagged or delayed effects of \( M_t \) and \( E_t \), our model implicitly assumes a lag structure with geometrically declining weights for \( M_t \) and \( E_t \) because the lagged endogenous variable appears as an explanatory variable in each equation. In this paper, we do not explore more complicated lag structures.

3. The Algorithms

The three algorithms presented in this paper are applicable to linear stochastic discrete-time econometric models with unknown parameters and additive Gaussian errors. We shall write the model
as a first-order linear difference equation

\begin{equation}
(3.1) \quad y_t^* = A y_{t-1}^* + C x_t + b_t + e_t,
\end{equation}

where \( A, C, \) and \( b_t \) are random parameters whose values will be estimated using the Bayesian techniques presented by Chow (1973a). The vector \( y_t^* \) is a stacked vector containing values of the endogenous variables and the instruments, \( x_t \) is a vector of instruments, \( b_t \) is a vector which models the effects of the noncontrollable exogenous variables, and \( e_t \) is a vector of random variables such that \( e_t \sim N(0, \Sigma) \). The \( e_t \) are assumed to be serially uncorrelated and uncorrelated with the random parameters \( A, C, \) and \( b_t \).

The objective of each of our three control algorithms is to minimize the expected value of the following quadratic welfare cost function

\begin{equation}
(3.2) \quad W = \frac{1}{2} \sum_{t=1}^{T} (y_t^* - a_t)' K_t (y_t^* - a_t),
\end{equation}

where \( T \) is the length of the planning horizon, \( a_t \) is the target value of \( y_t^* \), and \( K_t \) is a weighting matrix. Observe that (3.2) may be rewritten as

\begin{equation}
(3.3) \quad W = \frac{1}{2} \sum_{t=1}^{T} y_t^* K_t y_t^* + \sum_{t=1}^{T} y_t^* ' k_t + \text{constant},
\end{equation}

where \( k_t = -K_t a_t \). We solve this problem using the method of dynamic programming. Let \( E_{t-1} w_t \) denote the expected welfare cost from period \( t \) up to and including period \( T \), with the subscript \( t-1 \) indicating that the expectation is conditional on information
available at the end of period $t-1$. We first minimize $E_{T-1} w_T$ with respect to $x_T$. Letting $H_T = K_T$ and $h_T = k_T$, we obtain

$$E_{T-1} w_T = E_{T-1} \left( \frac{1}{2} y_T^* H_T y_T^* + y_T^* h_T \right) + \text{constant}'.
$$

Substituting (3.1) into (3.4) and partially differentiating with respect to $x_T$, we obtain the following feedback control equation, which yields the optimal value of $x_T$,

$$\hat{x}_T = G_T y_{T-1}^* + g_T,$$

where

$$G_T = -(E_{T-1} C' H_T C)^{-1} (E_{T-1} C' H_T A)$$

and

$$g_T = -(E_{T-1} C' H_T C)^{-1} [(E_{T-1} C' H_T b_T) + (E_{T-1} C') h_T].$$

Note that the feedback control equation is not linear in $y_{T-1}^*$ since the parameters $G_T$ and $g_T$ are functions of the posterior density of $A, C,$ and $b_T$ at the end of period $T-1$, which is a function of $y_{T-1}^*, y_{T-2}^*, \ldots$. After substituting $\hat{x}_T$ into (3.4) to obtain the optimal expected welfare cost $\hat{w}_T$ for the last period, we then approximate the function $\hat{w}_T$ by a modified second-order Taylor series expansion to obtain

$$w_T \approx \frac{1}{2} \sum_{t=1}^{T} y_t^* Q_T y_t^* + \sum_{t=1}^{T} y_t^* q_t + c_o .$$

Using Bellman’s principle of optimality, we minimize $E_{T-2} w_{T-1}$ with respect to $x_{T-1}$ under the assumption that the optimal value
of \( x_T \), i.e., \( \hat{x}_T \), will be selected in period \( T \). Hence, we seek to minimize

\[
(3.7) \quad E_{T-2}\hat{w}_{T-1} = E_{T-2}\left( \frac{1}{2}y_{T-1}^{*}\kappa_{T-1}y_{T-1}^{*} + y_{T-1}^{*}\kappa_{T-1} + \hat{w}_{T} \right) + \text{constant''}.
\]

Substituting the Taylor series approximation for \( \hat{w}_T \) and combining like terms in \( y_{T-1}^{*} \) within the expectation operator, we obtain

\[
(3.8) \quad E_{T-2}\hat{w}_{T-1} = E_{T-2}\left( \frac{1}{2}y_{T-1}^{*}h_{T-1}y_{T-1}^{*} + y_{T-1}^{*}h_{T-1} \right) + \text{constant '''},
\]

where \( h_{T-1} = \kappa_{T-1} + q_{T-1}^{T} \), \( h_{T-1} = \kappa_{T-1} + q_{T-1}^{T} \), and constant ''' absorbs those terms in (3.6) which are not dependent upon \( x_{T-1} \) and \( y_{T-1}^{*} \). It should be observed that (3.8) is identical in form to (3.4). Hence, we may solve for \( \hat{x}_{T-1} \) in the same manner that we solved for \( \hat{x}_T \). This backward induction procedure is repeated until we obtain values for \( \hat{x}_1 \) and \( \hat{w}_1 \). This algorithm, which we shall call method III, is a dual adaptive control algorithm. It gives only approximate solutions since it involves a quadratic approximation about a somewhat arbitrary tentative path.\(^3\)

In the certainty equivalence algorithm (method I), it is assumed that the values of the random parameters, \( A, C, \) and \( b_t \) are equal, with certainty, to their respective conditional expectations at time \( 0 \), i.e., \( \bar{A}, \bar{C}, \) and \( \bar{b}_t \). Hence, the parameters of the feedback control equations are

\[
G_T = -(\bar{C}'H_T\bar{C})^{-1}(\bar{C}'H_T\bar{A})
\]

and \( g_T = -(\bar{C}'H_T\bar{C})^{-1}[(\bar{C}'H_T\bar{B}) + \bar{C}'h_T] \). The feedback control equation is strictly linear in \( y_{T-1}^{*} \) and the function \( \hat{w}_T \) is truly
quadratic in \( y_{T-1}^* \). Therefore, there is no need to approximate 
\( \hat{w}_T \) by a second-order Taylor series expansion and thus the certainty 
equivalence solution to the control problem is exact.

Method II, which is another special case of method III, takes 
account of uncertainty in the parameters but does not anticipate 
future learning. In method II, all conditional expectations are 
evaluated at time 0 so that the coefficients of the feedback 
control equation are 
\[ g_T = -(E_0C'H_Tc)^{-1}(E_0C'H_Ta) \] 
and 
\[ g_T = -(E_0C'H_Tc)^{-1}[(E_0C'H_Tb_T)+(E_0C'h_T)] \]. As in method I, the 
parameters of the feedback control equation are independent of 
\( y_{T-1}^* \), and hence \( x_T \) is a linear function of \( y_{T-1}^* \). Thus \( \hat{w}_T \) 
is truly quadratic in \( y_{T-1}^* \) and method II yields an exact solu-
tion to the modified control problem.

4. **Policy Analysis Using the Simple Model**

Studies of the relative effectiveness of monetary and fiscal 
policy often focus on the size of long-run and short-run multipliers 
of the monetary and fiscal instruments, e.g., Kmenta and Smith 
(1973). However, Brainard (1967) argues that an examination of the 
minimum expected welfare cost attainable with a given set of instru-
ments is a more meaningful approach to the question of the effec-
tiveness of the instruments than is an examination of the multipliers 
of these instruments. Indeed, from the point of view of maximizing 
social welfare, any relevant features of the multipliers will be 
reflected in the minimum expected value of the welfare cost function 
and hence multiplier analysis is unnecessary, if not misleading. In this
paper we shall compare the minimum expected welfare cost attainable using only a discretionary monetary instrument with the minimum expected welfare cost attainable using only a discretionary fiscal instrument. However, before solving the control problem using only one discretionary instrument at a time, we solve the control problem using both $M_t$ and $E_t$ as discretionary instruments subject to feedback control. We rewrite the reduced form equations (3.4) and (3.5) as

$$y_t^* = Ay_{t-1}^* + Cx_t + b_t + e_t,$$

where

$$y_t^* = (C_t, I_t, E_t, M_t) ,$$

$$A = \begin{pmatrix} A_1 & 0 \\ 0 & 0 \end{pmatrix}$$

is a $4 \times 4$ matrix containing the $2 \times 2$ matrix $A_1$, $C = \begin{pmatrix} C_0 \\ I \end{pmatrix}$ is a $4 \times 2$ matrix containing the $2 \times 2$ matrix $C_0$, $x_t^* = (E_t, M_t)$ is the vector of instruments, $b_t = (b_t)$ is a 4-vector containing the 2-vector $b_t$, and $e_t = (e_t)$ is a 4-vector containing the 2-vector $e_t$.

Before proceeding with the application of the control algorithms, we must specify the following parameters of the welfare function: (1) $T$, the number of periods in the planning horizon; (2) $a_t$, the target values of $y_t^*$; and (3) $K_t$, the weighting matrices. We shall solve the control for a 6-period planning horizon, i.e., $T = 6$. In order to select appropriate target growth rates for $C_t$ and $I_t$, we examine the historical percentage growth rates shown in table 4.1.4. It should be noted that the growth rate for $I_t$ for the 11 quarters ending with 1963-IV is much higher than the growth
rate for the 40 quarters ending with 1963-IV because investment was near a cyclical low in 1961-II. With these historical growth rates in mind, we somewhat arbitrarily choose target growth rates of 1.25% per quarter for $C_t$ and $I_t$.

<table>
<thead>
<tr>
<th>Period</th>
<th>$C_t$</th>
<th>$I_t$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1954-I to 1963-IV</td>
<td>0.91</td>
<td>1.14</td>
</tr>
<tr>
<td>1961-I to 1963-IV</td>
<td>1.10</td>
<td>2.61</td>
</tr>
</tbody>
</table>

For the weighting matrices in the welfare function, we set $K_t = \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix}$, $t=1, \ldots, T$, where $I$ is the $2 \times 2$ identity matrix. Note that since the instruments are assigned zero weight in each $K_t$, they do not explicitly appear as arguments of the welfare function.

Since the ultimate objective of our analysis is to compare the relative effectiveness of monetary and fiscal policy in reaching given targets over time, we shall examine the welfare cost net of the costs directly associated with the instruments. We will, however, examine the stability of the instruments in each solution to make sure that they do not fluctuate excessively.

Method I is applied to the two-instrument, two-target control problem, with

$$\bar{A}_1 = \begin{pmatrix} 0.9266 & -0.0203 \\ 0.1527 & 0.3806 \end{pmatrix}, \quad b_t = \begin{pmatrix} -63.2386 \\ -210.8994 \end{pmatrix}$$

and

$$\bar{C}_0 = \begin{pmatrix} 0.3190 & 0.4206 \\ -0.0735 & 1.5389 \end{pmatrix}.$$
Then method II and method III are applied to the same problem. For
the quadratic approximation of method III, a deterministically
generated tentative path derived from the solution of method I is
used. The solutions to this problem by the three algorithms are
presented in Table 4.4.

In order to investigate the effectiveness of one instrument
alone, we shall assume that the instrument under consideration is
a discretionary instrument the values of which are chosen by the
policy maker subject to a feedback control. It is assumed that the
values of the other instrument are determined by a passive policy of
a constant percentage change per quarter. In the notation of (3.1),
the discretionary instrument is represented by the scalar \( x_t \), and
the passive instrument is modeled as a noncontrollable exogenous
variable which is absorbed in the value of \( b_t \). The solution to
the control problem will be sensitive to the values of \( b_t \),
\( t=1, \ldots, T \), and hence the values of the passive instrument must be
chosen judiciously. To determine the values of the passive policy
variable, we solve the six-period, two-instrument control problem
in which each of the instruments is constrained to change at a
constant rate throughout the planning horizon. The solution to this
problem, calculated under the assumption of certainty equivalence,
is shown in Table 4.2. There is no \textit{a priori} reason to believe that
the growth rate obtained in this manner for each instrument will be
optimal when the values of the other instrument are chosen subject
to feedback control. A better approach to selecting an optimal
TABLE 4.2 CERTAINTY EQUIVALENCE SOLUTION TO CONTROL PROBLEM WHEN BOTH INSTRUMENTS ARE PASSIVE

<table>
<thead>
<tr>
<th>Instrument</th>
<th>Period</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
<th>6</th>
<th>rate of change per period</th>
</tr>
</thead>
<tbody>
<tr>
<td>$E_t$ (billions)</td>
<td></td>
<td>110.9</td>
<td>112.4</td>
<td>113.9</td>
<td>115.4</td>
<td>116.9</td>
<td>118.4</td>
<td>+ 1.313%</td>
</tr>
<tr>
<td>$M_t$ (billions)</td>
<td></td>
<td>143.6</td>
<td>143.4</td>
<td>143.3</td>
<td>143.1</td>
<td>142.9</td>
<td>142.7</td>
<td>- 0.122%</td>
</tr>
</tbody>
</table>

NOTE: This solution was obtained using the OPTCDIAG option of the certainty equivalence program described in Douglas R. Chapman and Gregory C. Chow, "Optimal Control Programs: User's Guide," Econometric Research Program, Princeton University, Research Memorandum No. 181, May, 1972. Slight inconsistencies may appear above as a result of rounding since the program used a percentage growth rate with 6 decimal places. Also note that $\hat{\omega}_1 = 68,3074$.

growth rate for the passive instrument would be to solve the control problem repeatedly with one discretionary instrument and one passive instrument, allowing the growth rate for the passive instrument to vary in successive computations of the solution. The optimal growth rate for the passive instrument is the growth rate for which the optimal expected welfare cost is minimized. In lieu of performing an extensive search to determine the optimal growth rate for each instrument when the other instrument is discretionary, we merely examine two other growth rates for each instrument to check whether the growth rates shown in table 4.2 appear to be approximately optimal. The optimal welfare costs shown in table 4.3 were obtained from the solution, by method I, to the control problem in which the passive instrument grows at the given rate and the discretionary instrument is determined by
TABLE 4.3  OPTIMAL EXPECTED WELFARE COSTS (by Method I) 
WHEN ONE INSTRUMENT IS PASSIVE AND ONE INSTRUMENT IS DISCRETIONARY

<table>
<thead>
<tr>
<th>$E_t$ is passive</th>
<th>$M_t$ is passive</th>
</tr>
</thead>
<tbody>
<tr>
<td>Growth rate of $E_t$</td>
<td>Growth rate of $M_t$</td>
</tr>
<tr>
<td>$\hat{w}_1$</td>
<td>$\hat{w}_1$</td>
</tr>
<tr>
<td>+1.213%</td>
<td>48.7176</td>
</tr>
<tr>
<td>+1.313%</td>
<td>48.0508</td>
</tr>
<tr>
<td>+1.413%</td>
<td>48.1178</td>
</tr>
</tbody>
</table>

feedback control. Note that for each instrument, the value of $\hat{w}_1$ obtained using the growth rate from table 4.2 is smaller than the values of $\hat{w}_1$ obtained using growth rates 0.1% larger and 0.1% smaller than the growth rate in table 4.2. This result lends some credence to the assertion that, for each instrument, the growth rate in table 4.2 reasonably approximates the optimal rate when the other instrument is determined by feedback control.

The control problem is now solved using methods I, II, and III under the assumption that $E_t$ is a discretionary instrument and $M_t$ is a passive instrument exogenously set equal to the values given in table 4.2. This procedure is then repeated with $M_t$ as the discretionary instrument and $E_t$ as the passive instrument. The results of the control computations for period 1 are presented in table 4.4. Let $w_i(P)$ be the optimal expected welfare cost function from period 1 to period T where $i \in \{I, II, III\}$ refers to the algorithm employed and $P \subseteq P^* = \{E_t, M_t\}$
refers to the set of discretionary policy variables used in the application of the algorithm. Note that for each \( i \in \{I, II, III\} \),

\[
\min_{P \subseteq P^*} w_i(P) = w_i([E_t, M_t]),
\]

which is an illustration of the well-known fact that in a control problem with two targets, a lower optimal expected welfare cost is attainable using two instruments subject to feedback control than by using only one of these instruments subject to feedback control. More significant for our economic analysis, however, is the result that for each \( i \),

\[
w_i([E_t]) < w_i([M_t]).
\]

Therefore, assuming that the economy of the United States is appropriately modeled by (2.4) and (2.5), fiscal policy as represented by \( E_t \) is somewhat more effective with respect to the given welfare function than is monetary policy represented by \( M_t \). We note, however, that this difference is small, especially when we allow for uncertainty in method II.

To study our solution more closely, the value of \( w_i(P) \) can be decomposed into a deterministic welfare cost and a stochastic welfare cost. To compute the deterministic welfare cost, we first generate a deterministic time path for each of the endogenous variables by assuming that each parameter of the linear model (3.1) is equal to its point estimate at time 0, and that \( e_t = 0 \), for \( t = 1, \ldots, T \). The deterministic welfare cost is weighted sum of squared deviations of the deterministic time paths of the endogenous variables from their respective targets. The stochastic welfare cost is due to the randomness in \( y^*_t \) and results from the additive stochastic disturbance \( e_t \). Substituting the optimal value of
TABLE 4.4.  SOLUTIONS USING THE THREE CONTROL ALGORITHMS

<table>
<thead>
<tr>
<th>Instrument Set and Results</th>
<th>Method H₁,₁</th>
<th>II</th>
<th>III</th>
</tr>
</thead>
</table>
| \([E_1, M_1]\)             | \[
\begin{bmatrix}
1 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0
\end{bmatrix}
\] | \[
\begin{bmatrix}
1.4121 & -0.1265 & 0 & 0 \\
-0.1265 & 1.0789 & 0 & 0 \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0
\end{bmatrix}
\] | \[
\begin{bmatrix}
1.6397 & -0.1974 & -0.0184 & -0.0125 \\
-0.1974 & 1.1434 & 0.0040 & -0.0331 \\
-0.0184 & 0.0040 & 0.0062 & 0.0016 \\
-0.0125 & -0.0331 & 0.0016 & 0.0063
\end{bmatrix}
\] |
| \([G_1]\)                  | \[
\begin{bmatrix}
-2.6095 & 0.3666 & 0 & 0 \\
-0.2239 & -0.2298 & 0 & 0
\end{bmatrix}
\] | \[
\begin{bmatrix}
-1.6856 & 0.0647 & 0 & 0 \\
-0.2497 & -0.1997 & 0 & 0
\end{bmatrix}
\] | \[
\begin{bmatrix}
-1.7174 & 0.0765 & 0 & 0 \\
-0.2497 & -0.1991 & 0 & 0
\end{bmatrix}
\] |
| \([g'_1]\)                | \[
\begin{bmatrix}
1013.0050 & 243.1293 \\
709.0372 & 249.7009
\end{bmatrix}
\] | \[
\begin{bmatrix}
719.3068 & 249.0192
\end{bmatrix}
\] | \[
\begin{bmatrix}
111.7262 & 142.8996 \\
111.7831 & 142.8544
\end{bmatrix}
\] | \[
\begin{bmatrix}
111.6835 & 142.8729
\end{bmatrix}
\] |
| \([x'_1]\)                | \[
\begin{bmatrix}
111.7262 & 142.8996 \\
111.7831 & 142.8544
\end{bmatrix}
\] | \[
\begin{bmatrix}
111.6835 & 142.8729
\end{bmatrix}
\] | \[
\begin{bmatrix}
111.6835 & 142.8729
\end{bmatrix}
\] |
| \([\hat{w}_1]\)           | 35.5479     | 51.3291 | 48.3926 |

(continued)
<table>
<thead>
<tr>
<th>Instrument Set and Results</th>
<th>Method ( E_t )</th>
<th>Method ( M_t )</th>
</tr>
</thead>
<tbody>
<tr>
<td>I</td>
<td>II</td>
<td>III</td>
</tr>
<tr>
<td>( H_{1,1} )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1.1545 0.1586 0</td>
<td>1.7558 0.0296 0</td>
<td>2.3791 0.2670 -0.0309</td>
</tr>
<tr>
<td>0.1586 1.1628 0</td>
<td>0.0296 1.2169 0</td>
<td>0.2670 1.9123 0.0219</td>
</tr>
<tr>
<td>0 0 0</td>
<td>0 0 0</td>
<td>-0.0309 0.0219 0.0116</td>
</tr>
<tr>
<td>( G_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>[-2.7949 0.1763 0]</td>
<td>[-1.7012 0.0018 0]</td>
<td>[-1.6065 -0.0991 0]</td>
</tr>
<tr>
<td>( g_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>1094.6816</td>
<td>719.4268</td>
<td>694.4397</td>
</tr>
<tr>
<td>( x_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>110.3943</td>
<td>111.0514</td>
<td>111.1035</td>
</tr>
<tr>
<td>( w_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>44.1556</td>
<td>63.4141</td>
<td>74.2185</td>
</tr>
<tr>
<td>( H_{1,1} )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>3.0666 -0.2904 0</td>
<td>3.1216 -0.2876 0</td>
<td>7.0334 -1.1361 -0.0257</td>
</tr>
<tr>
<td>-0.2904 1.0408 0</td>
<td>-0.2876 1.1196 0</td>
<td>-1.1361 1.3879 -0.0329</td>
</tr>
<tr>
<td>0 0 0</td>
<td>0 0 0</td>
<td>-0.0257 -0.0329 0.0616</td>
</tr>
<tr>
<td>( G_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>[-0.3827 -0.2075 0]</td>
<td>[-0.3580 -0.1732 0]</td>
<td>[-0.4210 -0.1451 0]</td>
</tr>
<tr>
<td>( g_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>298.0639</td>
<td>286.1260</td>
<td>306.2156</td>
</tr>
<tr>
<td>( x_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>142.9770</td>
<td>142.8952</td>
<td>142.8845</td>
</tr>
<tr>
<td>( w_1 )</td>
<td></td>
<td></td>
</tr>
<tr>
<td>48.0508</td>
<td>64.7409</td>
<td>77.2147</td>
</tr>
</tbody>
</table>
\( x_t \) from (3.5) into (4.1), we obtain
\[
(4.2) \quad y_t^* = (A + CG_t)y_{t-1}^* + Cg_t + b_t + e_t
\]
Assuming that the covariance matrix of \( e_t \) is \( \Sigma_e \) for \( t > 0 \), it can be shown that the covariance matrix of \( y_t^* \) is given by the recursive formula
\[
(4.3) \quad \Sigma_{y_t}^* = (A + CG_t) \Sigma_{y_t}^*(A + CG_t)' + \Sigma_e
\]
The stochastic welfare cost is equal to
\[
(4.4) \quad w_s = \frac{1}{2} \sum_{t=1}^{T} \text{tr}(K_t V_t)
\]
where \( V_t \) is the estimate of \( \Sigma_{y_t}^* \) based on the estimate of \( \Sigma_e \) at the current time.

In figure 4.1 we present the target time path for \( C_t \) and the deterministic time paths for \( C_t \) using the instrument sets \( \{E_t\} \) and \( \{M_t\} \). In addition to the deterministic time path for each instrument set, we present the values of \( C_t \) one standard deviation above and below the deterministic time path of \( C_t \). For each time period, the vertical distance between the deterministic time path of \( C_t \) for a given instrument set and the target time path of \( C_t \) essentially measures the square root of the deterministic cost attributable to \( C_t \). Similarly, for each period, the standard deviation of \( C_t \) around its deterministic time path reflects the stochastic cost attributable to \( C_t \) in that period.

Note that the deterministic time path of \( C_t \) for the instrument set \( \{E_t\} \) is generally above the target time path.
Figure 4.1. Expected Time Paths of $C_t$ with Standard Deviation Bands (Obtained from Certainty Equivalence Solutions)
Figure 4.2  Expected time paths of $I_t$ with standard deviation bands (obtained from certainty equivalence solutions)
whereas for the instrument set \( \{M_t\} \) it is generally below the
target time path. If the deterministic cost comprised a major
portion of the expected welfare cost, we might have to consider
the given quadratic welfare cost function to be inappropriate
because it assigns costs to the overachievement of the targets for
\( C_t \) through the use of the instrument set \( \{E_t\} \) as well as to the
underachievement of the targets for \( C_t \) through the use of \( \{M_t\} \).
However, we note that for each instrument set, the standard de-
VIation of the stochastic variation around the deterministic time
path of \( C_t \) far outveighs the deterministic "standard deviation"
of the deterministic time path around the target time path. Hence,
the adverse effects of assigning deterministic costs to expected
positive deviations from the target values may be neglected since
they appear to be unimportant. We also note that the standard
deVIation band around the deterministic time path obtained using
\( \{E_t\} \) lies within the standard deviation band around the determin-
istic time path obtained using \( \{M_t\} \), except for period 1. Hence,
the stochastic cost attributable to \( C_t \) is smaller for \( \{E_t\} \) than
for \( \{M_t\} \). Figure 4.2 is analogous to 4.1 except that the endo-
genous target variable is \( I_t \). As in Figure 4.1, we observe that
the stochastic welfare cost is much larger than the deterministic
welfare cost.

Table 4.5 summarizes the results presented in Figures
4.1 and 4.2. In this table, the deterministic welfare cost is
expressed as an average over the six-period planning horizon. The
stochastic welfare cost for each target variable is the average variance of that variable around its deterministic time path. The last two columns of table 4.5 present the square roots of the corresponding values in the first two columns of the table and represent deviations in terms of 1958 dollars. It is clear from Table 4.5 that the total cost attributable to $I_t$ is greater than the total cost attributable to $C_t$, and the stochastic cost is much greater than the deterministic cost for each instrument set.

Since the instruments receive zero weights in the welfare function, we shall briefly examine the dynamic characteristics of the time paths of the instruments (derived from method I) to determine whether they are highly volatile. We note that for the instrument set $\{E_t, M_t\}$, the period-to-period fluctuations of $E_t$ along its deterministic time path are all less than 1.3 billion 1958 dollars, and for $\{E_t\}$ all of the deterministic changes are less than 2.5 billion 1958 dollars. Furthermore for each instrument set, the standard deviation of $E_t$ around its deterministic time path remains fairly stable and is less than 5.3 billion 1958 dollars in each period. For each of the instrument sets $\{E_t, M_t\}$ and $\{M_t\}$, the deterministic period-to-period fluctuations of $M_t$ are all less than 0.1 billion 1958 dollars. The standard deviation of $M_t$ remains fairly stable at about 0.9 billion 1958 dollars for $\{E_t, M_t\}$ and about 1.1 billion 1958 dollars for $\{M_t\}$. Hence, it appears that for each instrument, neither the deterministic period-to-period fluctuations nor the stochastic variation around the
\[\text{TABLE 4.5. AVERAGE WELFARE COSTS PER PERIOD}\]

<table>
<thead>
<tr>
<th>Instrument Set</th>
<th>Target</th>
<th>Variance (Welfare Cost) [billions of 1958 dollars]²</th>
<th>Standard Deviations [billions of 1958 dollars]</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Deterministic</td>
<td>Stochastic</td>
</tr>
<tr>
<td>({E_t})</td>
<td>(C_t)</td>
<td>0.028</td>
<td>3.816</td>
</tr>
<tr>
<td></td>
<td>(I_t)</td>
<td>0.696</td>
<td>10.178</td>
</tr>
<tr>
<td></td>
<td>Total</td>
<td>0.724</td>
<td>13.994</td>
</tr>
<tr>
<td>({M_t})</td>
<td>(C_t)</td>
<td>0.072</td>
<td>7.273</td>
</tr>
<tr>
<td></td>
<td>(I_t)</td>
<td>0.016</td>
<td>8.656</td>
</tr>
<tr>
<td></td>
<td>Total</td>
<td>0.088</td>
<td>15.929</td>
</tr>
<tr>
<td>({E_t, M_t})</td>
<td>(C_t)</td>
<td>0</td>
<td>3.775</td>
</tr>
<tr>
<td></td>
<td>(I_t)</td>
<td>0</td>
<td>8.074</td>
</tr>
<tr>
<td></td>
<td>Total</td>
<td>0</td>
<td>11.849</td>
</tr>
</tbody>
</table>

**NOTE:** Costs exclude factor of \(1/2\) which appears in (2.4) and (4.6).

Deterministic time path is large enough to present serious problems of implementation.

We observe in table 4.4 that for a given instrument set, the coefficients of the feedback control equation are subject to considerable variation across algorithms with the introduction of uncertainty and the anticipation of learning. However, it should be noted that the optimal values of the instrument do not appear to be very sensitive to the presence of uncertainty or to the anticipation of learning. In table 4.6, we present the percentage
### Table 4.6. Percentage Variation in $\hat{x}_1$ Across the Three Algorithms

<table>
<thead>
<tr>
<th>Instrument Set</th>
<th>Instrument</th>
<th>% Variation Across Algorithms</th>
</tr>
</thead>
<tbody>
<tr>
<td>{E_t, M_t}</td>
<td>E_t</td>
<td>.09%</td>
</tr>
<tr>
<td></td>
<td>M_t</td>
<td>.03%</td>
</tr>
<tr>
<td>{E_t}</td>
<td>E_t</td>
<td>.64</td>
</tr>
<tr>
<td>{M_t}</td>
<td>M_t</td>
<td>.06</td>
</tr>
</tbody>
</table>

**NOTE:** The percentage variation is the ratio of the range of $\hat{x}_1$ to the value of $\hat{x}_1$ obtained for method I.

Variation across the three algorithms of the optimal first-period settings of the instruments for each of the three control problems. Note that when the policy maker treats both $E_t$ and $M_t$ as discretionary instruments, there is an extremely small percentage variation in $\hat{x}_1$ across the three algorithms. This result suggests that for the purpose of determining the optimal values of $E_1$ and $M_1$, it makes little difference whether the effects of uncertainty and learning are considered.

### 5. Concluding Remarks

Using the very simple macro-econometric model presented in Section 2, we found that fiscal policy, represented by $E_t$, is more effective than monetary policy, represented by $M_t$, with
respect to the given welfare function. Note, however, that this result does not imply that the policy maker should treat $M_t$ as a passive instrument not subject to feedback control. The results presented in table 4.4 indicate that the minimum expected welfare cost is significantly lower when the policy maker selects the values of both $E_t$ and $M_t$ subject to feedback control than when $E_t$ is the only discretionary instrument. We also observed that the values of the instruments required to achieve the minimum expected welfare cost appear to be free from wild fluctuations over time and do not thereby present a difficult problem of implementation.

In the evaluation of the relative effectiveness of the monetary and fiscal instruments, we allowed for uncertainty and the possibility of learning in the computation of the optimal control solutions and the associated welfare losses. By examining the effectiveness of policy within the framework of the three different algorithms and their different assumptions regarding uncertainty and learning, our analysis has a broader basis than if we had used only method I with its restrictive assumptions of certainty equivalence. We noted that although the introduction of uncertainty may significantly change the coefficients of the feedback control equation, the optimal first-period policy is rather insensitive to uncertainty in the parameters of the linear model. The implication of this result for policy formulation is that we may fairly accurately determine the optimal values of $E_1$ and $M_1$ by any of the three algorithms discussed in this paper.
FOOTNOTES

*I would like to express my most sincere thanks to Professor Gregory C. Chow, my thesis advisor, for generously sharing his time and his ideas with me.


2. In the calculations summarized later we employ a standard Taylor series with cross-partial terms of the form \((y_t^* - a_t)^tK_{s,s}(y_s^* - a_s)\). However, we let \(K_{t,s} = 0\) for all \(t,s\).

3. In our calculations, we obtain the tentative path by using the certainty equivalence algorithm to determine \(\hat{x}_t\), and then apply the estimated model, without random disturbance, to \(\hat{x}_t\), to generate \(y_t^*\), for \(t = 1, \ldots, T - 1\).

4. Since \(E_t\) and \(M_t\) have zero weight in the welfare function, it is not necessary to specify targets for these variables.

5. This expected welfare cost ignores the factor of \(\frac{1}{2}\) in (3.2).


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